



## Research Article

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# ***p*-variation and Chung's LIL of sub-bifractional Brownian motion and applications**

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**Abstract:** Let  $S^{H,K} = \{S^{H,K}(t), t \geq 0\}$  be the sub-bifractional Brownian motion, with  $H \in (0, 1)$  and  $K \in (0, 1]$ . We investigate its  $p$ -variation and Chung's law of the iterated logarithm. In addition, we give some applications of these properties.

**Keywords:** sub-bifractional Brownian motion,  $p$ -variation, Chung's law of the iterated logarithm

**MSC 2020:** 60G15, 60G18, 60F25

## 1 Introduction

El-Nouty and Journé [1] introduced the process  $S^{H,K} = \{S^{H,K}(t), t \geq 0\}$  with  $H \in (0, 1)$  and  $K \in (0, 1]$ , named the sub-bifractional Brownian motion (sbfBm) and defined by

$$S^{H,K}(t) = \frac{1}{2^{(2-K)/2}}(B^{H,K}(t) + B^{H,K}(-t)),$$

where  $\{B^{H,K}(t), t \in \mathbb{R}\}$  is a bifractional Brownian motion (bfBm) with  $H \in (0, 1)$  and  $K \in (0, 1]$ . Clearly, the sbfBm is a centered Gaussian process such that  $S^{H,K}(0) = 0$ , with probability 1, and  $\text{Var}(S^{H,K}(t)) = (2^K - 2^{2HK-1})t^{2HK}$ . Note that  $(2H-1)K - 1 < K - 1 \leq 0$ , we have  $2HK - 1 < K$ . We can prove that  $S^{H,K}$  is self-similar with index  $HK$ . When  $K = 1$ ,  $S^{H,1}$  is the subfractional Brownian motion (sfBm). We can easily obtain that for all  $s, t \geq 0$ ,

$$\mathbb{E}(S^{H,K}(t)S^{H,K}(s)) = (t^{2H} + s^{2H})^K - \frac{1}{2}(t + s)^{2HK} - \frac{1}{2}|t - s|^{2HK} \quad (1.1)$$

and

$$C_1|t - s|^{2HK} \leq \mathbb{E}[(S^{H,K}(t) - S^{H,K}(s))^2] \leq C_2|t - s|^{2HK}, \quad (1.2)$$

where

$$C_1 = \min\{2^K - 1, 2^K - 2^{2HK-1}\}, \quad C_2 = \max\{1, 2 - 2^{2HK-1}\}.$$

(See [1]).

El-Nouty and Journé [1] proved that the sbfBm is a quasi-helix in the sense of Kahane, and the upper classes of some of its increments are characterized by an integral test. Kuang [2] investigated the collision local time of two independent sbfBms. Kuang and Li [3] obtained Berry-Esséen bounds and proved the almost sure

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central limit theorem for the quadratic variation in the sbfBm. Finally, Kuang and Xie [4] studied least squares-type estimators for the drift parameters in the sub-bifractional Vasicek processes.

In this article, we investigate  $p$ -variation and Chung's law of the iterated logarithm (Chung's LIL) of sbfBm. In addition, we give some applications of these properties.

Throughout this article, some specific constants in Section  $i$  are numbered as  $c_{i,1}, c_{i,2}, \dots$ .

This study is organized as follows: In Section 2, we study  $p$ -variation. Section 3 is devoted to Chung's LIL. Section 4 contains some applications of its properties.

## 2 $p$ -variation

The variation in Gaussian processes was studied extensively since the works of [5], which proved almost sure convergence to 1 of the quadratic variation  $\sum_{j=1}^{2^n} |B(j/2^n) - B((j-1)/2^n)|^2$  of the Brownian motion  $B$  on  $[0, 1]$ . Many new results about the variation in Gaussian processes with stationary increments were obtained (refer [6–9] and references therein). Wang [10] studied the  $p$ -variation in bfBm. Shen et al. [11] obtained the power variation in the sfBm.

We will consider  $p$ -variation in sbfBm by using the ideas of Wang [10] and Shen et al. [11]. However, the increments of sbfBm are not independent and not stationary, this causes some difficulties to investigate the variation in the process. In order to overcome the difficulties, we develop a stochastic integral representation of sbfBm.

Now, we state our main results in this section as follows.

**Theorem 2.1.** *Let  $T > 0$ ,  $a > 0$ , and  $v_n = n^a$ . Then, for any  $p \geq 1$ , we have, as  $n \rightarrow \infty$ ,*

$$\frac{1}{v_n^{1-pHK}} \sum_{j=1}^{[Tv_n]} \left| S^{H,K} \left( \frac{j}{v_n} \right) - S^{H,K} \left( \frac{j-1}{v_n} \right) \right|^p \rightarrow T \frac{2^{p/2} \Gamma(\frac{p+1}{2})}{\Gamma(\frac{1}{2})}, \quad a.s., \quad (2.1)$$

where  $[x]$  denotes the integer part of  $x > 0$ , and  $\Gamma(x) = \int_0^\infty t^{x-1} e^{-t} dt$  for  $x > 0$ , which is a Gamma function.

**Corollary 2.2.** *Let  $T > 0$ ,  $a > 0$ , and  $v_n = n^a$ . Then, for any  $p \geq 1$ , we have, as  $n \rightarrow \infty$ ,*

$$\frac{1}{v_n^{1-pHK}} \sum_{j=1}^{[Tv_n]} \left| \left( S^{H,K} \left( \frac{j}{v_n} \right) \right)^2 - \left( S^{H,K} \left( \frac{j-1}{v_n} \right) \right)^2 \right|^p \rightarrow T \frac{2^{(3p)/2} \Gamma(\frac{p+1}{2})}{\Gamma(\frac{1}{2})} \int_0^T |S^{H,K}(x)|^p dx, \quad a.s. \quad (2.2)$$

**Theorem 2.3.** *Let  $T > 0$ ,  $a > 0$ , and  $v_n = n^a$ . Then, we have, as  $n \rightarrow \infty$ ,*

$$\frac{1}{v_n^{1-HK}} \sum_{j=1}^{[Tv_n]} \sup_{\frac{j-1}{v_n} < t, u < \frac{j}{v_n}} |S^{H,K}(t) - S^{H,K}(u)| \rightarrow 2E \left( \sup_{0 \leq t \leq T^{1/(HK)}} S^{H,K}(t) \right), \quad a.s. \quad (2.3)$$

In order to prove Theorems 2.1 and 2.3, we give some technical lemmas. Lemma 2.1 is a Fernique-type inequality for  $S^{H,K}$ .

**Lemma 2.1.** *For any  $\varepsilon > 0$ , there exists a positive constant  $c_{2,1} = c_{2,1}(\varepsilon) > 0$  such that*

$$P \left\{ \sup_{0 \leq t \leq T} \sup_{0 \leq s \leq a} |S^{H,K}(t+s) - S^{H,K}(t)| \geq x a^{HK} \right\} \leq c_{2,1} \left( \frac{T}{a} + 1 \right) e^{-\frac{x^2}{2(1+\varepsilon)}}, \quad (2.4)$$

for any  $T \geq 0$ ,  $a > 0$ , and  $x \geq x_0 > 0$  with some  $x_0 > 0$ .

**Proof.** By (1.2) and the inequality for the normal distribution function  $\Phi(x) : 1 - \Phi(x) \leq \frac{1}{x}e^{-\frac{x^2}{2}}$  for all  $x > 0$ , we obtain that

$$\mathbf{P}\{|S^{H,K}(t+h) - S^{H,K}(t)| \geq xh^{HK}\} \leq c_{2,2}e^{-\frac{x^2}{2}}$$

for any  $t \geq 0$ ,  $h > 0$ , and  $x \geq x^* > 0$  with some  $x^* > 0$ . Therefore, by Lemmas 2.1 and 2.2 in [8] (when applied to  $\sigma_1(h) = h^{HK}$  and  $\sigma_2(\cdot) \equiv 0$ ), we obtain (2.4) immediately.  $\square$

Lemma 2.2 is from [12].

**Lemma 2.2.** *Let  $X = \{X(t), t \in \mathbf{R}\}$  be a centered Gaussian process in  $\mathbf{R}$  and let  $F \subset \mathbf{R}$  be a closed set equipped with the canonical metric defined by*

$$d(s, t) = [\mathbf{E}(X(s) - X(t))^2]^{1/2}. \quad (2.5)$$

*Then, there exists a positive constant  $c_{2,3}$  such that for all  $u > 0$ ,*

$$\mathbf{P}\left\{\sup_{s,t \in F}|X(s) - X(t)| \geq c_{2,3}\left(u + \int_0^D \sqrt{\log N_d(F, \varepsilon)} d\varepsilon\right)\right\} \leq \exp\left(-\frac{u^2}{D^2}\right), \quad (2.6)$$

*where  $N_d(F, \varepsilon)$  denotes the smallest number of open  $d$ -balls of radius  $\varepsilon$  needed to cover  $F$  and where  $D = \sup\{d(s, t) : s, t \in F\}$  is the diameter of  $F$ .*

**Lemma 2.3.** *If  $X(t)$  and  $Y(t)$  are a.s. bounded, centered Gaussian processes on  $\Lambda$  such that  $\mathbf{E}(X^2(t)) = \mathbf{E}(Y^2(t))$  for all  $t \in \Lambda$ , and*

$$\mathbf{E}[(X(t) - X(s))^2] \leq \mathbf{E}[(Y(t) - Y(s))^2], \quad \forall s, t \in \Lambda,$$

*then for all real  $\lambda$ ,*

$$\mathbf{P}\left\{\sup_{t \in \Lambda} X(t) > \lambda\right\} \leq \mathbf{P}\left\{\sup_{t \in \Lambda} Y(t) > \lambda\right\}$$

*and*

$$\mathbf{E}\left\{\sup_{t \in \Lambda} X(t)\right\} \leq \mathbf{E}\left\{\sup_{t \in \Lambda} Y(t)\right\}.$$

**Proof.** It is Slepian's inequality (see, p. 49 in [13]).  $\square$

In order to solve the dependence structure of  $S^{H,K}$  and to create independence, we will develop the stochastic integral representation of  $S^{H,K}$ . By Lamperti's transformation [14], we define Gaussian process  $Y = \{Y(t), t \in \mathbf{R}\}$  as follows:

$$Y(t) = e^{-HKt}S^{H,K}(e^t), \quad t \in \mathbf{R}. \quad (2.7)$$

The covariance function  $r(t) := \mathbf{E}(Y(0)Y(t))$  is given by

$$\begin{aligned} r(t) &= e^{-HKt}\left\{(1 + e^{2Ht})^K - \frac{1}{2}(1 + e^t)^{2HK} - \frac{1}{2}|1 - e^t|^{2HK}\right\} \\ &= e^{HKt}\left\{(1 + e^{-2Ht})^K - \frac{1}{2}(1 + e^{-t})^{2HK} - \frac{1}{2}|1 - e^{-t}|^{2HK}\right\} \\ &= r(-t). \end{aligned} \quad (2.8)$$

Hence,  $r(t)$  is an even function and, by (2.8) and the Taylor expansion, we verify that  $r(t) = O(e^{-\beta t})$  as  $t \rightarrow \infty$ , where  $\beta = H(2 - K)$ . It follows that  $r(\cdot) \in L^1(\mathbf{R})$ . By (2.8) and the Taylor expansion we obtain

$$r(t) \sim 2^K - 2^{2HK-1} - \frac{1}{2} |t|^{2HK}, \quad t \rightarrow 0. \quad (2.9)$$

By Bochner's theorem [15],  $Y$  has the stochastic integral representation:

$$Y(t) = \int_{\mathbf{R}} e^{i\lambda t} W(d\lambda), \quad \forall t \in \mathbf{R}, \quad (2.10)$$

where  $W$  is a complex Gaussian measure with control measure  $\Delta$ , whose Fourier transform is  $r(\cdot)$ . The measure  $\Delta$  is called the spectral measure of  $Y$ .

Since  $r(\cdot) \in L^1(\mathbf{R})$ , the spectral measure  $\Delta$  of  $Y$  has a continuous density function  $f(\lambda)$ , which can be represented as the inverse Fourier transform of  $r(\cdot)$

$$f(\lambda) = \frac{1}{\pi} \int_0^\infty r(t) \cos(t\lambda) dt. \quad (2.11)$$

Similar to the proof of (2.10) in [16], we can obtain

$$f(\lambda) \sim c_{2,4} |\lambda|^{-(1+2HK)}, \quad \text{as } \lambda \rightarrow \infty, \quad (2.12)$$

where  $c_{2,4} > 0$  is an explicit constant depending only on  $HK$ .

By (2.7) and (2.10), we obtain

$$S^{H,K}(t) = t^{HK} \int_{\mathbf{R}} e^{i\lambda \log t} W(d\lambda), \quad \forall t > 0. \quad (2.13)$$

We list two properties of the spectral density  $f(\lambda)$  of  $Y$ . They follow from (2.12), or from (2.9) and the truncation inequalities in [17], page 209, refer also [18].

**Lemma 2.4.** *There exist positive constants  $c_{2,5}$  and  $c_{2,6}$  such that for  $u > 1$ ,*

$$\int_{|\lambda| < u} \lambda^2 f(\lambda) d\lambda \leq c_{2,5} u^{2(1-HK)} \quad (2.14)$$

and

$$\int_{|\lambda| \geq u} f(\lambda) d\lambda \leq c_{2,6} u^{-2HK}. \quad (2.15)$$

**Proof of Theorem 2.1.** Without loss of generality, we suppose  $Tv_h$  is an integer. For integers  $n$  and  $j \geq 1$ , we take  $a_{n,j} = (jn)^\beta$ , where  $\beta > 0$  is a constant. Define two Gaussian processes

$$X_{n,j}^{(1)}(t) = t^{HK} \int_{|\lambda| \in (a_{n,j}, a_{n,j+1}]} e^{i\lambda \log t} W(d\lambda)$$

and

$$X_{n,j}^{(2)}(t) = t^{HK} \int_{|\lambda| \notin (a_{n,j}, a_{n,j+1}]} e^{i\lambda \log t} W(d\lambda).$$

Clearly, by (2.13), we have

$$S^{H,K}(t) = X_{n,j}^{(1)}(t) + X_{n,j}^{(2)}(t), \quad \text{for all } t \geq 0. \quad (2.16)$$

It is important to note that for a fixed  $n$ , the Gaussian processes  $X_{n,j}^{(1)}(t), j = 1, 2, \dots$ , are independent; moreover, for every  $j \geq 1$ ,  $X_{n,j}^{(1)}(t)$  and  $X_{n,j}^{(2)}(t)$  are also independent. Since

$$\begin{aligned}
& \left| \frac{1}{v_n^{1-pHK}} \sum_{j=1}^{Tv_n} \left| S^{H,K} \left( \frac{j}{v_n} \right) - S^{H,K} \left( \frac{j-1}{v_n} \right) \right|^p - T \frac{2^{p/2} \Gamma(\frac{p+1}{2})}{\Gamma(\frac{1}{2})} \right| \\
& \leq \frac{1}{v_n} \sum_{j=1}^{Tv_n} v_n^{pHK} \left\| S^{H,K} \left( \frac{j}{v_n} \right) - S^{H,K} \left( \frac{j-1}{v_n} \right) \right\|^p - \left\| X_{n,j}^{(1)} \left( \frac{j}{v_n} \right) - X_{n,j}^{(1)} \left( \frac{j-1}{v_n} \right) \right\|^p \\
& \quad + \frac{1}{v_n} \sum_{j=1}^{Tv_n} v_n^{pHK} \left\| \mathbf{E} \left| S^{H,K} \left( \frac{j}{v_n} \right) - S^{H,K} \left( \frac{j-1}{v_n} \right) \right|^p - \mathbf{E} \left| X_{n,j}^{(1)} \left( \frac{j}{v_n} \right) - X_{n,j}^{(1)} \left( \frac{j-1}{v_n} \right) \right|^p \right\|^p \\
& \quad + \frac{1}{v_n} \sum_{j=1}^{Tv_n} v_n^{pHK} \left\| X_{n,j}^{(1)} \left( \frac{j}{v_n} \right) - X_{n,j}^{(1)} \left( \frac{j-1}{v_n} \right) \right\|^p - \mathbf{E} \left\| X_{n,j}^{(1)} \left( \frac{j}{v_n} \right) - X_{n,j}^{(1)} \left( \frac{j-1}{v_n} \right) \right\|^p \\
& \quad + \left| \frac{1}{v_n} \sum_{j=1}^{Tv_n} v_n^{pHK} \mathbf{E} \left| S^{H,K} \left( \frac{j}{v_n} \right) - S^{H,K} \left( \frac{j-1}{v_n} \right) \right|^p - T \frac{2^{p/2} \Gamma(\frac{p+1}{2})}{\Gamma(\frac{1}{2})} \right| \\
& = I_1 + I_2 + I_3 + I_4.
\end{aligned}$$

In the following, we will show that the terms  $I_1$  and  $I_3$  almost surely converge to zero,  $I_2$  and  $I_4$  converge to zero, as  $n \rightarrow \infty$ , respectively.

First, we prove for  $a > 0$ ,  $T > 0$ , and  $p \geq 1$ , as  $n \rightarrow \infty$ ,

$$I_1 = \frac{1}{v_n} \sum_{j=1}^{Tv_n} v_n^{pHK} \left\| S^{H,K} \left( \frac{j}{v_n} \right) - S^{H,K} \left( \frac{j-1}{v_n} \right) \right\|^p - \left\| X_{n,j}^{(1)} \left( \frac{j}{v_n} \right) - X_{n,j}^{(1)} \left( \frac{j-1}{v_n} \right) \right\|^p \rightarrow 0, \quad \text{a.s.} \quad (2.17)$$

In fact,

$$\begin{aligned}
& v_n^{pHK} \left\| S^{H,K} \left( \frac{j}{v_n} \right) - S^{H,K} \left( \frac{j-1}{v_n} \right) \right\|^p - \left\| X_{n,j}^{(1)} \left( \frac{j}{v_n} \right) - X_{n,j}^{(1)} \left( \frac{j-1}{v_n} \right) \right\|^p \\
& \leq c_{2,7} \left( \sup_{0 \leq t \leq T} |S^{H,K}(t)|^{p-1} + \sup_{0 \leq t \leq T} |X_{n,j}^{(1)}(t)|^{p-1} \right) v_n^{(p-1)HK} Y_{n,j},
\end{aligned} \quad (2.18)$$

where

$$Y_{n,j} = v_n^{HK} \sup_{\frac{j-1}{v_n} < t, u < \frac{j}{v_n}} |X_{n,j}^{(2)}(t) - X_{n,j}^{(2)}(u)|, \quad (2.19)$$

and where we use the fact

$$||x|^p - |y|^p| \leq p2^{p-1}||x|^{p-1} + |y|^{p-1}||x - y|.$$

We have

$$\begin{aligned}
& v_n^{pHK} \left\| \mathbf{E} \left| S^{H,K} \left( \frac{j}{v_n} \right) - S^{H,K} \left( \frac{j-1}{v_n} \right) \right|^p - \mathbf{E} \left\| X_{n,j}^{(1)} \left( \frac{j}{v_n} \right) - X_{n,j}^{(1)} \left( \frac{j-1}{v_n} \right) \right\|^p \right\| \\
& \leq c_{2,8} \left( \mathbf{E} \sup_{0 \leq t \leq T} |S^{H,K}(t)|^{2(p-1)} + \mathbf{E} \sup_{0 \leq t \leq T} |X_{n,j}^{(1)}(t)|^{2(p-1)} \right)^{\frac{1}{2}} v_n^{(p-1)HK} (\mathbf{E} Y_{n,j}^2)^{\frac{1}{2}}.
\end{aligned} \quad (2.20)$$

For  $Y_{n,j}$ , by Lemmas 2.2 and 2.4, elementary calculus can show that there exists  $n_0$  such that for any  $n \geq n_0$ , for every  $1 \leq j \leq Tv_n$  and for any  $t > 0$ ,

$$\mathbf{P}(Y_{n,j} > t) \leq c_{2,9} n^{a+\beta} \exp(-c_{2,10} n^{2\beta HK - 2a} t^2). \quad (2.21)$$

Thus, for any  $\varepsilon > 0$ , we obtain

$$\begin{aligned} \mathbf{P}\left(\frac{1}{v_n} \sum_{j=1}^{Tv_n} v_n^{(p-1)HK} Y_{n,j} > \varepsilon\right) &\leq \mathbf{P}\left(\max_{1 \leq j \leq Tv_n} Y_{n,j} > \frac{\varepsilon}{T} v_n^{-(p-1)HK}\right) \\ &\leq \sum_{j=1}^{Tv_n} \mathbf{P}\left(Y_{n,j} > \frac{\varepsilon}{T} v_n^{-(p-1)HK}\right) \\ &\leq c_{2,11} n^{2a+\beta} \exp(-c_{2,12} n^{2\beta HK - 2a - 2a(p-1)HK}). \end{aligned}$$

Taking  $\beta > 0$  large enough such that  $2\beta HK - 2a - 2a(p-1)HK > 0$ , by the Borel-Cantelli lemma, we have

$$\lim_{n \rightarrow \infty} \frac{1}{v_n} \sum_{j=1}^{Tv_n} v_n^{(p-1)HK} Y_{n,j} = 0, \quad \text{a.s.} \quad (2.22)$$

Combining (2.18) and (2.22), we prove that (2.17) holds.

Second, we prove for  $a > 0$ ,  $T > 0$ , and  $p \geq 1$ , as  $n \rightarrow \infty$ ,

$$I_2 = \frac{1}{v_n} \sum_{j=1}^{Tv_n} v_n^{pHK} \left| \mathbf{E} \left| S^{H,K} \left( \frac{j}{v_n} \right) - S^{H,K} \left( \frac{j-1}{v_n} \right) \right|^p - \mathbf{E} \left| X_{n,j}^{(1)} \left( \frac{j}{v_n} \right) - X_{n,j}^{(1)} \left( \frac{j-1}{v_n} \right) \right|^p \right| \rightarrow 0. \quad (2.23)$$

In fact, for any  $1 \leq j \leq Tv_n$  and  $r > 2$ , by (2.21) and Hölder's inequality, we obtain

$$\begin{aligned} (\mathbf{E}(Y_{n,j})^2)^{\frac{1}{2}} &\leq (\mathbf{E}(Y_{n,j})^r)^{\frac{1}{r}} \\ &= \left( \int_0^\infty \mathbf{P}((Y_{n,j})^r > s) ds \right)^{\frac{1}{r}} \\ &\leq \left( c_{2,9} n^{a+\beta} \int_0^\infty \exp\left(-c_{2,10} n^{2\beta HK - 2a} s^{\frac{2}{r}}\right) ds \right)^{\frac{1}{r}} \\ &= c_{2,13} n^{-\beta HK + a + \frac{a+\beta}{r}}, \end{aligned} \quad (2.24)$$

where we use by letting  $s^{\frac{2}{r}} = t$ , then

$$\begin{aligned} \int_0^\infty \exp\left(-c_{2,10} n^{2\beta HK - 2a} s^{\frac{2}{r}}\right) ds &= \frac{r}{2} \int_0^\infty t^{\frac{r}{2}-1} \exp(-c_{2,10} n^{2\beta HK - 2a} t) dt \\ &= \frac{r}{2} \int_0^\infty \left( \frac{y}{c_{2,10} n^{2\beta HK - 2a}} \right)^{\frac{r}{2}-1} \cdot \frac{e^{-y} dy}{c_{2,10} n^{2\beta HK - 2a}} \\ &= \frac{\frac{r}{2} \Gamma(\frac{r}{2}) n^{(-\beta HK + a)r}}{(c_{2,10})^{\frac{r}{2}}}. \end{aligned}$$

Hence,

$$\frac{1}{v_n} \sum_{j=1}^{Tv_n} v_n^{(p-1)HK} (\mathbf{E}(Y_{n,j})^2)^{\frac{1}{2}} \leq c_{2,14} n^{a(p-1)HK - \beta HK + a + \frac{a+\beta}{r}}.$$

Taking first  $\beta > 0$  large enough and then taking  $r > 2$  large enough such that  $a(p-1)HK - \beta HK + a + \frac{a+\beta}{r} < 0$ . Therefore, we obtain

$$\lim_{n \rightarrow \infty} \frac{1}{v_n} \sum_{j=1}^{Tv_n} v_n^{(p-1)HK} (\mathbf{E}(Y_{n,j})^2)^{\frac{1}{2}} = 0. \quad (2.25)$$

Similar to (2.24), by using (2.4), for any  $p > 1$ , we have that

$$\mathbf{E} \left( \sup_{0 \leq t \leq T} |S^{H,K}(t)|^{2(p-1)} \right) \leq c_{2,15}. \quad (2.26)$$

Since, for any  $j \geq 1$ ,  $t \geq 0$ ,  $h > 0$ , we have

$$\mathbf{E}(X_{n,j}^{(1)}(t+h) - X_{n,j}^{(1)}(t))^2 \leq \mathbf{E}(S^{H,K}(t+h) - S^{H,K}(t))^2.$$

Then, (2.4) remains true for  $X_{n,j}^{(1)}$ . Thus, similar to (2.26), we obtain for any  $1 \leq j \leq Tv_n$ ,

$$\mathbf{E} \left( \sup_{0 \leq t \leq T} |X_{n,j}^{(1)}(t)|^{2(p-1)} \right) \leq c_{2,16}.$$

Hence, combining (2.20) and (2.25), we have that (2.23) holds.

Third, we prove that for  $a > 0$ ,  $T > 0$ , and  $p \geq 1$ , as  $n \rightarrow \infty$ ,

$$I_3 = \frac{1}{v_n} \sum_{j=1}^{Tv_n} v_n^{pHK} \left| X_{n,j}^{(1)} \left( \frac{j}{v_n} \right) - X_{n,j}^{(1)} \left( \frac{j-1}{v_n} \right) \right|^p - \mathbf{E} \left| X_{n,j}^{(1)} \left( \frac{j}{v_n} \right) - X_{n,j}^{(1)} \left( \frac{j-1}{v_n} \right) \right|^p \rightarrow 0, \quad \text{a.s.} \quad (2.27)$$

In fact, since for a fixed  $n$ , the processes  $\{X_{n,j}^{(1)}(t), t \geq 0\}, j = 1, 2, \dots, Tv_n$  are independent and so are  $\left| X_{n,j}^{(1)} \left( \frac{j}{v_n} \right) - X_{n,j}^{(1)} \left( \frac{j-1}{v_n} \right) \right|^p, j = 1, 2, \dots, Tv_n$ . For any  $\varepsilon > 0$  and  $r > 1$ , by Markov inequality and the moment inequality of partial sums of independent random variables, we have

$$\begin{aligned} & \mathbf{P} \left( \sum_{j=1}^{Tv_n} v_n^{pHK} \left| X_{n,j}^{(1)} \left( \frac{j}{v_n} \right) - X_{n,j}^{(1)} \left( \frac{j-1}{v_n} \right) \right|^p - \mathbf{E} \left| X_{n,j}^{(1)} \left( \frac{j}{v_n} \right) - X_{n,j}^{(1)} \left( \frac{j-1}{v_n} \right) \right|^p \right) > \varepsilon v_n \right) \\ & \leq \frac{c_{2,17}}{v_n^r} \mathbf{E} \left[ \left| \sum_{j=1}^{Tv_n} v_n^{pHK} \left| X_{n,j}^{(1)} \left( \frac{j}{v_n} \right) - X_{n,j}^{(1)} \left( \frac{j-1}{v_n} \right) \right|^p - \mathbf{E} \left| X_{n,j}^{(1)} \left( \frac{j}{v_n} \right) - X_{n,j}^{(1)} \left( \frac{j-1}{v_n} \right) \right|^p \right|^{pr} \right] \\ & \leq \frac{c_{2,18}}{v_n^r} v_n^{\frac{r}{2}+1} \sum_{j=1}^{Tv_n} \mathbf{E} \left[ v_n^{prHK} \left| X_{n,j}^{(1)} \left( \frac{j}{v_n} \right) - X_{n,j}^{(1)} \left( \frac{j-1}{v_n} \right) \right|^p - \mathbf{E} \left| X_{n,j}^{(1)} \left( \frac{j}{v_n} \right) - X_{n,j}^{(1)} \left( \frac{j-1}{v_n} \right) \right|^p \right]^{pr} \\ & \leq \frac{c_{2,19}}{v_n^{\frac{r}{2}+1}} \sum_{j=1}^{Tv_n} \mathbf{E} \left[ \left| v_n^{pHK} \left| X_{n,j}^{(1)} \left( \frac{j}{v_n} \right) - X_{n,j}^{(1)} \left( \frac{j-1}{v_n} \right) \right|^p \right|^{pr} \right]. \end{aligned} \quad (2.28)$$

Since

$$\sigma_{n,j}^2 = \mathbf{E} \left| X_{n,j}^{(1)} \left( \frac{j+t}{v_n} \right) - X_{n,j}^{(1)} \left( \frac{j+u}{v_n} \right) \right|^2 \leq \mathbf{E} \left| S^{H,K} \left( \frac{j+t}{v_n} \right) - S^{H,K} \left( \frac{j+u}{v_n} \right) \right|^2$$

and

$$\begin{aligned} & \frac{\mathbf{E} \left( S^{H,K} \left( \frac{j+t}{v_n} \right) - S^{H,K} \left( \frac{j+u}{v_n} \right) \right)^2}{\frac{1}{v_n^{2HK}}} \\ & = (2^K - 2^{2HK-1})(j+t)^{2HK} + (2^K - 2^{2HK-1})(j+u)^{2HK} - 2((j+t)^{2H} + (j+u)^{2H})^K \\ & \quad + (2j+t+u)^{2HK} + |t-u|^{2HK} \\ & \rightarrow |t-u|^{2HK}, \quad \text{as } j \rightarrow \infty, \end{aligned} \quad (2.29)$$

for any  $t, u \in [0, 1]$  and  $n \geq 1$ .

Hence, for  $t = 1, u = 0$ , there exists  $j_0 \geq 1$  such that for any  $j \geq j_0$ ,

$$\mathbf{E} \left[ \left| v_n^{pHK} \left| X_{n,j}^{(1)} \left( \frac{j}{v_n} \right) - X_{n,j}^{(1)} \left( \frac{j-1}{v_n} \right) \right|^p \right|^{pr} \right] \leq c_{2,20} v_n^{prHK} \sigma_{n,j}^{pr} \leq c_{2,21},$$

where we have used the fact: let  $X$  be a random variable following an  $N(0, \sigma^2)$ , then for any  $\gamma > 0$ ,

$$\mathbf{E}(|X|^\gamma) = \frac{2^{\frac{\gamma}{2}} \Gamma\left(\frac{\gamma+1}{2}\right)}{\Gamma\left(\frac{1}{2}\right)} \sigma^\gamma. \quad (2.30)$$

Therefore,

$$\mathbf{P}\left(\sum_{j=1}^{Tv_n} v_n^{pHK} \left|X_{n,j}^{(1)}\left(\frac{j}{v_n}\right) - X_{n,j}^{(1)}\left(\frac{j-1}{v_n}\right)\right|^p - \mathbf{E} \left|X_{n,j}^{(1)}\left(\frac{j}{v_n}\right) - X_{n,j}^{(1)}\left(\frac{j-1}{v_n}\right)\right|^p\right) > \varepsilon v_n \leq c_{2,22} n^{-\frac{ar}{2}}.$$

Taking  $r > 1$  large enough such that  $\frac{ar}{2} > 1$  and by Borel-Cantelli lemma, we obtain (2.27) holds.

Finally, we prove that for  $a > 0$ ,  $T > 0$ , and  $p \geq 1$ , as  $n \rightarrow \infty$ ,

$$I_4 = \frac{1}{v_n} \sum_{j=1}^{Tv_n} v_n^{pHK} \mathbf{E} \left|S^{H,K}\left(\frac{j}{v_n}\right) - S^{H,K}\left(\frac{j-1}{v_n}\right)\right|^p - T \frac{2^{p/2} \Gamma(\frac{p+1}{2})}{\Gamma(\frac{1}{2})} \rightarrow 0. \quad (2.31)$$

In fact, by (2.29) and (2.30), we have for large  $j$ ,

$$v_n^{pHK} \mathbf{E} \left|S^{H,K}\left(\frac{j}{v_n}\right) - S^{H,K}\left(\frac{j-1}{v_n}\right)\right|^p \rightarrow \frac{2^{p/2} \Gamma(\frac{p+1}{2})}{\Gamma(\frac{1}{2})}.$$

Hence, (2.31) holds. Thus, the proof of Theorem 2.1 is complete.  $\square$

**Proof of Corollary 2.2.** By Theorem 2.1, following the same lines as the proof of Theorem 1.2 in [19], we can easily prove the corollary, and omit the details.  $\square$

**Proof of Theorem 2.3.** For simplicity, we assume that  $Tv_n$  is an integer. For  $a > 0$ , we denote

$$\begin{aligned} \xi_{n,j} &= \xi_{n,j}(S^{H,K}, a) = \sup_{\frac{j-1}{v_n} < t, u < \frac{j}{v_n}} |S^{H,K}(t) - S^{H,K}(u)|, \\ \eta_{n,j} &= \eta_{n,j}(S^{H,K}, a) = v_n^{HK} \xi_{n,j}. \end{aligned}$$

We first prove that for every  $a > 0$ ,

$$\lim_{n \rightarrow \infty} \frac{1}{v_n^{1-HK}} \sum_{j=1}^{Tv_n} |\xi_{n,j} - \mathbf{E} \xi_{n,j}| = 0, \quad \text{a.s.} \quad (2.32)$$

Denote  $\zeta_{n,j} = \xi_{n,j}(X_{n,j}^{(1)}, a)$ ,  $Y_{n,j} = v_n^{HK} \xi_{n,j}(X_{n,j}^{(2)}, a)$  ( $Y_{n,j}$  is actually defined by (2.19)). In order to show (2.32), it is enough to prove that

$$\lim_{n \rightarrow \infty} \frac{1}{v_n^{1-HK}} \sum_{j=1}^{Tv_n} |\zeta_{n,j} - \mathbf{E} \zeta_{n,j}| = 0, \quad \text{a.s.}, \quad (2.33)$$

$$\lim_{n \rightarrow \infty} \frac{1}{v_n} \sum_{j=1}^{Tv_n} |Y_{n,j}| = 0, \quad \text{a.s.}, \quad (2.34)$$

and

$$\lim_{n \rightarrow \infty} \frac{1}{v_n^{1-HK}} \sum_{j=1}^{Tv_n} |\mathbf{E} \zeta_{n,j} - \mathbf{E} \xi_{n,j}| = 0, \quad (2.35)$$

By equalities (2.22) and (2.23), we can obtain equalities (2.34) and (2.35), respectively. We are preparing to prove (2.33).

In fact, since for a fixed  $n$ ,  $X_{n,j}^{(1)}$ ,  $j = 1, 2, \dots, Tn$ , are independent; so are  $v_n^{HK} \zeta_{n,j}$ ,  $j = 1, 2, \dots, Tn$ , similar to (2.28), for any  $\varepsilon > 0$  and  $r > 1$ , we have

$$\mathbf{P}\left(\sum_{j=1}^{Tn} |v_n^{HK} \zeta_{n,j} - \mathbf{E}(v_n^{HK} \zeta_{n,j})| > \varepsilon v_n\right) \leq \frac{c_{2,23}}{v_n^r} \sum_{j=1}^{Tn} \mathbf{E}[(v_n^{HK} \zeta_{n,j})^r]. \quad (2.36)$$

By Lemma 2.1, we obtain for every  $t > t_0$  with some  $t_0 > 0$  and  $1 \leq j \leq Tn$ ,

$$\mathbf{P}(\eta_{n,j} > t) = \mathbf{P}\left(\sup_{\frac{j-1}{v_n} < t, u < \frac{j}{v_n}} |S^{H,K}(t) - S^{H,K}(u)| \geq tv_n^{-HK}\right) \leq c_{2,24} v_n e^{-c_{2,25} t^2}.$$

Hence, by (2.16) and (2.21), we obtain

$$\mathbf{P}(v_n^{HK} \zeta_{n,j} > t) \leq \mathbf{P}\left(\eta_{n,j} > \frac{t}{2}\right) + \mathbf{P}\left(Y_{n,j} > \frac{t}{2}\right) \leq c_{2,26} n^{a+\beta} e^{-c_{2,27} t^2}.$$

Therefore, for every  $1 \leq j \leq Tn$ ,

$$\begin{aligned} \mathbf{E}(v_n^{HK} \zeta_{n,j})^r &= \int_0^\infty \mathbf{P}\left(v_n^{HK} \zeta_{n,j} > t^{\frac{1}{r}}\right) dt \\ &\leq t_0 + \int_{t_0}^\infty \mathbf{P}\left(v_n^{HK} \zeta_{n,j} > t^{\frac{1}{r}}\right) dt \\ &\leq t_0 + c_{2,28} n^{a+\beta} \int_0^\infty \exp(-c_{2,29} t^{2/r}) dt \leq c_{2,30} n^{a+\beta}. \end{aligned}$$

Hence, by (2.36), we obtain

$$\mathbf{P}\left(\sum_{j=1}^{Tn} |v_n^{HK} \zeta_{n,j} - \mathbf{E}(v_n^{HK} \zeta_{n,j})| > \varepsilon v_n\right) \leq \frac{c_{2,31} n^{a+\beta}}{v_n^r} = c_{2,32} n^{-\frac{ar}{2} + a + \beta}.$$

Taking  $r > 1$  large enough such that  $-\frac{ar}{2} + a + \beta < -1$ , by the Borel-Cantelli lemma, we have

$$\frac{1}{v_n^{1-HK}} \sum_{j=1}^{Tn} |\zeta_{n,j} - \mathbf{E}\zeta_{n,j}| \rightarrow 0, \quad \text{a.s.},$$

as  $n \rightarrow \infty$ . Thus, (2.33) holds.

In order to finish the proof of Theorem 2.3, by the self-similarity of  $S^{H,K}$ , we only need to show

$$\lim_{n \rightarrow \infty} \frac{1}{v_n} \sum_{j=1}^{Tn} \mathbf{E}(\eta_{n,j}) = 2T \mathbf{E}\left(\sup_{0 \leq t \leq 1} S^{H,K}(t)\right). \quad (2.37)$$

By (2.29), there exists  $j_0 > 0$  such that for every  $t, u \in [0, 1]$  and all  $j \geq j_0$ ,

$$\lim_{n \rightarrow \infty} \frac{\mathbf{E}\left(S^{H,K}\left(\frac{j+t}{v_n}\right) - S^{H,K}\left(\frac{j+u}{v_n}\right)\right)^2}{\frac{|t-u|^{2HK}}{v_n^{2HK}}} = 1.$$

Hence, by Lemma 2.3, for every  $j \geq j_0$ ,

$$\lim_{n \rightarrow \infty} \mathbf{E}\left[\sup_{0 < t < 1} \left(S^{H,K}\left(\frac{j-1+t}{v_n}\right) - S^{H,K}\left(\frac{j-1}{v_n}\right)\right)\right] = \lim_{n \rightarrow \infty} \mathbf{E}\left[\sup_{0 < t < 1} \left(S^{H,K}\left(\frac{t}{v_n}\right)\right)\right].$$

Therefore,

$$\begin{aligned}
\lim_{n \rightarrow \infty} \mathbf{E}(\xi_{n,j}) &= \lim_{n \rightarrow \infty} \mathbf{E} \left[ \sup_{0 < t, u < 1} \left| \left( S^{H,K} \left( \frac{j-1+t}{v_n} \right) - S^{H,K} \left( \frac{j-1}{v_n} \right) \right) - \left( S^{H,K} \left( \frac{j-1+u}{v_n} \right) - S^{H,K} \left( \frac{j-1}{v_n} \right) \right) \right| \right] \\
&= \lim_{n \rightarrow \infty} \mathbf{E} \left[ \sup_{0 < t, u < 1} \left| \left( S^{H,K} \left( \frac{j-1+t}{v_n} \right) - S^{H,K} \left( \frac{j-1}{v_n} \right) \right) - \left( S^{H,K} \left( \frac{j-1+u}{v_n} \right) - S^{H,K} \left( \frac{j-1}{v_n} \right) \right) \right| \right] \\
&= \lim_{n \rightarrow \infty} 2 \mathbf{E} \left[ \sup_{0 < t < 1} \left| S^{H,K} \left( \frac{j-1+t}{v_n} \right) - S^{H,K} \left( \frac{j-1}{v_n} \right) \right| \right] \\
&= \lim_{n \rightarrow \infty} 2 \mathbf{E} \left[ \sup_{0 < t < 1} \left| S^{H,K} \left( \frac{t}{v_n} \right) \right| \right].
\end{aligned}$$

Hence, by the self-similarity of  $S^{H,K}$ , for any  $j \geq j_0$ ,

$$\begin{aligned}
\lim_{n \rightarrow \infty} \mathbf{E}(\eta_{n,j}) &= \lim_{n \rightarrow \infty} \mathbf{E}(v_n^{HK} \xi_{n,j}) \\
&= 2 \lim_{n \rightarrow \infty} \left( v_n^{HK} \mathbf{E} \left[ \sup_{0 < t < 1} \left| S^{H,K} \left( \frac{t}{v_n} \right) \right| \right] \right) \\
&= 2 \mathbf{E} \left[ \sup_{0 < t < 1} (S^{H,K}(t)) \right].
\end{aligned}$$

By (2.26), we have

$$\begin{aligned}
\max_{1 \leq j \leq j_0} \mathbf{E}(\eta_{n,j}) &\leq \mathbf{E} \left[ v_n^{HK} \sup_{0 < t, u < \frac{j_0}{v_n}} |S^{H,K}(t) - S^{H,K}(u)| \right] \\
&= 2 \mathbf{E} \left[ v_n^{HK} \sup_{0 < t < \frac{j_0}{v_n}} S^{H,K}(t) \right] \\
&= 2 j_0^{HK} \mathbf{E} \left[ \sup_{0 < t < 1} S^{H,K}(t) \right] \\
&\leq c_{2,33}.
\end{aligned}$$

Therefore,

$$\lim_{n \rightarrow \infty} \frac{1}{v_n} \sum_{j=1}^{Tv_n} \mathbf{E}(\eta_{n,j}) = \lim_{n \rightarrow \infty} \frac{1}{v_n} \sum_{j=1}^{j_0} \mathbf{E}(\eta_{n,j}) + \lim_{n \rightarrow \infty} \frac{1}{v_n} \sum_{j=j_0+1}^{Tv_n} \mathbf{E}(\eta_{n,j}) = 2T \mathbf{E} \left[ \sup_{0 < t < 1} (S^{H,K}(t)) \right].$$

The proof of Theorem 2.3 is completed.  $\square$

### 3 Chung's LIL

In [16,18,20,21] the authors established Chung's LIL for fBm and other strongly locally nondeterministic Gaussian processes with stationary increments. Luan [22] obtained Chung's LIL for sbfBm. In this section, we prove the Chung's LIL for sbfBm  $S^{H,K}$  in  $\mathbf{R}$ .

**Theorem 3.1.** *Let  $S^{H,K} = \{S^{H,K}(t), t \geq 0\}$  be the sbfBms in  $\mathbf{R}$ , with  $H \in (0, 1)$  and  $K \in (0, 1]$ . Then, there exists a positive and finite constant  $c_{3,1}$  such that*

$$\liminf_{r \rightarrow 0} \frac{\max_{t \in [0, r]} |S^{H,K}(t)|}{r^{HK}/(\log \log(1/r))^{HK}} = c_{3,1}, \quad \text{a.s.} \tag{3.1}$$

In order to prove Theorem 3.1, we need several lemmas. Lemma 3.1 shows that the sbfBms  $S^{H,K}$  has strong local nondeterminism. Lemma 3.2 gives estimates on the small ball probability of  $S^{H,K}$ .

**Lemma 3.1.** *For all constants  $0 < a < b$ ,  $S^{H,K}$  is strongly locally  $\varphi$ -nondeterministic on  $I = [a, b]$  with  $\varphi(r) = r^{2HK}$ . That is, there exist positive constants  $c_{3,2}$  and  $r_0$  such that for all  $t \in I$  and all  $0 < r \leq \min\{t, r_0\}$ ,*

$$\text{Var}\{S^{H,K}(t)|S^{H,K}(s) : s \in I, r \leq |s - t| \leq r_0\} \geq c_{3,2}\varphi(r). \quad (3.2)$$

**Proof.** See the proof of Proposition 2.1 in [2], the proof follows the same line as Proposition 2.1 in [16].  $\square$

**Lemma 3.2.** *There exist positive constants  $c_{3,3}$  and  $c_{3,4}$  such that for all  $t_0 \in [0, 1]$  and  $x \in (0, 1)$ ,*

$$\exp\left(-\frac{c_{3,3}}{x^{1/(HK)}}\right) \leq \mathbf{P}\left[\max_{t \in [0, 1]} |S^{H,K}(t) - S^{H,K}(t_0)| \leq x\right] \leq \exp\left(-\frac{c_{3,4}}{x^{1/(HK)}}\right). \quad (3.3)$$

**Proof.** By Lemma 3.1 and (1.2), we know that  $S^{H,K}$  satisfies conditions (C1) and (C2) of [23]. Hence, this lemma holds by Theorem 3.1 of [23].  $\square$

The following Lemma 3.3 is from [16], which provides a zero-one law for ergodic self-similar processes.

**Lemma 3.3.** *Let  $X = \{X_t, t \in \mathbf{R}\}$  be a separable, self-similar process with index  $k$ . We assume that  $X_0 = 0$  and that  $X$  is ergodic. Then, for any increasing function  $\psi : \mathbf{R}_+ \rightarrow \mathbf{R}_+$ , we have  $\mathbf{P}(E_{k,\psi}) = 0$  or 1, where*

$$E_{k,\psi} = \left\{ \omega : \text{there exists } \delta > 0 \text{ such that } \sup_{0 \leq s \leq t} |X_s| \geq t^k \psi(t) \text{ for all } 0 < t \leq \delta \right\}.$$

By a result of [24] on ergodicity and mixing properties of stationary Gaussian processes, we see that  $S^{H,K}$  is mixing. Hence, we can obtain the following lemma.

**Lemma 3.4.** *There exists a constant  $c_{3,5} \in [0, \infty]$  such that*

$$\liminf_{t \rightarrow 0_+} \frac{(\log \log(1/t))^{HK}}{t^{HK}} \max_{0 \leq s \leq t} |S^{H,K}(s)| = c_{3,5}, \quad a.s. \quad (3.4)$$

**Proof.** We take  $\psi_c(t) = c(\log \log(1/t))^{-HK}$  and define  $c_{3,5} = \sup\{c \geq 0 : \mathbf{P}(E_{k,\psi_c}) = 1\}$ . Then, (3.4) holds from Lemma 3.3.  $\square$

Theorem 3.1 will be established if we prove  $c_{3,5} \in (0, \infty)$  from Lemma 3.4. This is where Lemmas 3.2 and 2.2 are needed.

Now, we proceed to prove Theorem 3.1.

**Proof of Theorem 3.1.** We prove the lower bound first. For any integer  $n \geq 1$ , let  $r_n = e^{-n}$ . Let  $0 < \gamma < c_{3,4}$  be a constant and consider the event

$$A_n = \left\{ \max_{0 \leq s \leq r_n} |S^{H,K}(s)| \leq \gamma^{HK} r_n^{HK} / (\log \log(1/r_n))^{HK} \right\}. \quad (3.5)$$

Then, by the self-similarity of  $S^{H,K}$  and Lemma 3.2,

$$\begin{aligned} \mathbf{P}\{A_n\} &= \mathbf{P}\left\{ \max_{0 \leq s \leq r_n} |S^{H,K}(s)| \leq \gamma^{HK} r_n^{HK} / (\log \log(1/r_n))^{HK} \right\} \\ &= \mathbf{P}\left\{ \max_{0 \leq s \leq 1} |S^{H,K}(r_n s)| \leq \gamma^{HK} r_n^{HK} / (\log \log(1/r_n))^{HK} \right\} \end{aligned} \quad (3.6)$$

$$\begin{aligned}
&= \mathbf{P} \left\{ r_n^{HK} \max_{0 \leq s \leq 1} |S^{H,K}(s)| \leq \gamma^{HK} r_n^{HK} / (\log \log(1/r_n))^{HK} \right\} \\
&= \mathbf{P} \left\{ \max_{0 \leq s \leq 1} |S^{H,K}(s)| \leq \gamma^{HK} / (\log \log(1/r_n))^{HK} \right\} \\
&\leq \exp \left( -\frac{c_{3,4}}{\gamma} \log n \right) \\
&= n^{-c_{3,4}/\gamma}.
\end{aligned} \tag{3.6}$$

Since  $\sum_{n=1}^{\infty} \mathbf{P}\{A_n\} < \infty$ , by the Borel-Cantelli lemma, we obtain

$$\liminf_{n \rightarrow \infty} \frac{\max_{0 \leq s \leq r_n} |S^{H,K}(s)|}{r_n^{HK} / (\log \log(1/r_n))^{HK}} \geq c_{3,4} \quad \text{a.s.} \tag{3.7}$$

By (3.7) and a standard monotonicity argument, we have

$$\liminf_{r \rightarrow 0} \frac{\max_{0 \leq s \leq r} |S^{H,K}(s)|}{r^{HK} / (\log \log(1/r))^{HK}} \geq c_{3,6} \quad \text{a.s.} \tag{3.8}$$

We will prove the upper bound by the following stochastic integral representation of  $S^{H,K}$ . For every  $t > 0$ , by (2.13), we have

$$S^{H,K}(t) = t^{HK} \int_{\mathbb{R}} e^{i\lambda \log t} W(d\lambda).$$

For every integer  $n \geq 1$ , we take

$$t_n = n^{-n} \quad \text{and} \quad d_n = n^{\beta}, \tag{3.9}$$

where  $\beta > 0$  is a constant whose value will be determined later. It is sufficient to prove that there exists a finite constant  $c_{3,7}$  such that

$$\liminf_{n \rightarrow \infty} \frac{\max_{0 \leq s \leq t_n} |S^{H,K}(s)|}{t_n^{HK} / (\log \log(1/t_n))^{HK}} \leq c_{3,7} \quad \text{a.s.} \tag{3.10}$$

Define two Gaussian processes,  $X_n^1$  and  $X_n^2$ , by

$$X_n^1(t) = t^{HK} \int_{|\lambda| \in (d_{n-1}, d_n]} e^{i\lambda \log t} W(d\lambda) \tag{3.11}$$

and

$$X_n^2(t) = t^{HK} \int_{|\lambda| \notin (d_{n-1}, d_n]} e^{i\lambda \log t} W(d\lambda), \tag{3.12}$$

respectively. Clearly,  $S^{H,K}(t) = X_n^1(t) + X_n^2(t)$  for all  $t \geq 0$ . It is important to note that the Gaussian processes  $X_n^1 (n = 1, 2, \dots)$  are independent; moreover, for every  $n \geq 1$ , the processes  $X_n^1$  and  $X_n^2$  are also independent.

Let  $h(r) = r^{HK} (\log \log(1/r))^{-HK}$ . We make the following two claims:

(i) There is a constant  $\gamma > 0$  such that

$$\sum_{n=1}^{\infty} \mathbf{P} \left\{ \max_{s \in [0, t_n]} |X_n^1(s)| \leq \gamma^{HK} h(t_n) \right\} = \infty. \tag{3.13}$$

(ii) For every  $\varepsilon > 0$ ,

$$\sum_{n=1}^{\infty} \mathbf{P} \left\{ \max_{s \in [0, t_n]} |X_n^2(s)| > \varepsilon h(t_n) \right\} < \infty. \tag{3.14}$$

Since the events in (3.13) are independent, we see that (3.10) follows from (3.13), (3.14), and a standard Borel-Cantelli argument.

It remains to verify the claims (i) and (ii) above. By Lemma 3.2 and Anderson's inequality [25], we obtain

$$\begin{aligned}
 \mathbf{P} \left\{ \max_{s \in [0, t_n]} |X_n^1(s)| \leq \gamma^{HK} h(t_n) \right\} &\geq \mathbf{P} \left\{ \max_{s \in [0, t_n]} |S^{H,K}(s)| \leq \gamma^{HK} h(t_n) \right\} \\
 &= \mathbf{P} \left\{ \max_{s \in [0, 1]} |S^{H,K}(t_n s)| \leq \gamma^{HK} h(t_n) \right\} \\
 &= \mathbf{P} \left\{ t_n^{HK} \max_{s \in [0, 1]} |S^{H,K}(s)| \leq \gamma^{HK} h(t_n) \right\} \\
 &= \mathbf{P} \left\{ \max_{s \in [0, 1]} |S^{H,K}(s)| \leq \gamma^{HK} / (\log \log(1/t_n))^{HK} \right\} \\
 &\geq \exp \left( - \frac{c_{3,3}}{\gamma} \log(n \log n) \right) \\
 &= (n \log n)^{-c_{3,3}/\gamma}.
 \end{aligned} \tag{3.15}$$

Thus, (i) holds for  $\gamma \geq c_{3,3}$ .

In order to prove (ii), we divide  $[0, t_n]$  into  $p_n + 1$  non-overlapping subintervals  $J_{n,j} = [a_{n,j-1}, a_{n,j}]$ ,  $j = 0, 1, \dots, p_n$ , and then apply Lemma 2.2 to  $X_n^2$  on each of  $J_{n,j}$ . Let  $\beta > 0$  be the constant in (3.9) and take  $J_{n,0} = [0, t_n n^{-\beta}]$ . After  $J_{n,0}$  has been defined, we take  $a_{n,j+1} = a_{n,j}(1 + n^{-\beta})$ . It can be verified that the number of such subintervals of  $[0, t_n]$  satisfies the following bound:

$$p_n + 1 \leq cn^\beta \log n. \tag{3.16}$$

Moreover, for every  $j \geq 1$ , if  $s, t \in J_{n,j}$  and  $s < t$ , then we have  $t/s - 1 \leq n^{-\beta}$  and this yields

$$t - s \leq sn^{-\beta} \quad \text{and} \quad \log \left( \frac{t}{s} \right) \leq n^{-\beta}. \tag{3.17}$$

(1.2) implies that the canonical metric  $d$  for the process  $X_n^2$  satisfies

$$d(s, t) \leq c |s - t|^{HK} \quad \text{for all } s, t > 0 \tag{3.18}$$

and  $d(0, s) \leq ct_n^{HK} n^{-\beta HK}$  for every  $s \in J_{n,0}$ . It follows that  $D_0 = \sup\{d(s, t); s, t \in J_{n,0}\} \leq ct_n^{HK} n^{-\beta HK}$ , and

$$N_d(J_{n,0}, \varepsilon) \leq \frac{t_n n^{-\beta}}{(\varepsilon/c)^{1/(HK)}}. \tag{3.19}$$

Some simple calculations yield

$$\begin{aligned}
 \int_0^{D_0} \sqrt{\log N_d(J_{n,0}, \varepsilon)} d\varepsilon &\leq \int_0^{ct_n^{HK} n^{-\beta HK}} \sqrt{\log \left( \frac{t_n n^{-\beta}}{(\varepsilon/c)^{1/(HK)}} \right)} d\varepsilon \\
 &= \int_0^1 ct_n^{HK} n^{-\beta HK} \sqrt{\log \left( \frac{1}{u} \right)^{1/(HK)}} du \\
 &= c \sqrt{1/(HK)} t_n^{HK} n^{-\beta HK} \int_0^1 \sqrt{\log \left( \frac{1}{u} \right)} du \\
 &= c_{3,8} t_n^{HK} n^{-\beta HK}.
 \end{aligned} \tag{3.20}$$

It follows from Lemma 2.2 and (3.20) that

$$\begin{aligned}
\mathbf{P} \left\{ \max_{s \in J_{n,0}} |X_n^2(s)| > \varepsilon h(t_n) \right\} &= \mathbf{P} \left\{ \max_{s \in J_{n,0}} |X_n^2(s) - X_n^2(0)| > \varepsilon h(t_n) \right\} \\
&\leq \mathbf{P} \left\{ \max_{s,t \in J_{n,0}} |X_n^2(s) - X_n^2(t)| > c_{3,9} 2u \right\} \\
&\leq \mathbf{P} \left\{ \max_{s,t \in J_{n,0}} |X_n^2(s) - X_n^2(t)| > c_{3,9} \left( u + \int_0^{D_0} \sqrt{\log N_d(J_{n,0}, \varepsilon)} d\varepsilon \right) \right\} \\
&\leq \exp \left( -\frac{u^2}{(D_0)^2} \right) \\
&\leq \exp \left( -c \frac{t_n^{2HK} (\log(n \log n))^{-2HK}}{t_n^{2HK} n^{-2\beta HK}} \right) \\
&= \exp \left( -c \frac{n^{2\beta HK}}{(\log(n \log n))^{2HK}} \right),
\end{aligned} \tag{3.21}$$

where  $u = \frac{\varepsilon}{2c_{3,9}} h(t_n)$ , which is larger than  $\int_0^{D_0} \sqrt{\log N_d(J_{n,0}, \varepsilon)} d\varepsilon$ .

For every  $1 \leq j \leq p_n$ , we estimate the  $d$ -diameter of  $J_{n,j}$ . It follows from (3.12) that for any  $s, t \in J_{n,j}$  with  $s < t$ ,

$$\begin{aligned}
\mathbf{E}(X_n^2(s) - X_n^2(t))^2 &= \int_{|\lambda| \leq d_{n-1}} |t^{HK} e^{i\lambda \log t} - s^{HK} e^{i\lambda \log s}|^2 f(\lambda) d\lambda \\
&\quad + \int_{|\lambda| > d_n} |t^{HK} e^{i\lambda \log t} - s^{HK} e^{i\lambda \log s}|^2 f(\lambda) d\lambda \\
&=: T_1 + T_2.
\end{aligned} \tag{3.22}$$

For  $T_2$ , we have, for all  $s, t \in J_{n,j}$ ,

$$T_2 \leq 4t_n^{2HK} \int_{|\lambda| > d_n} f(\lambda) d\lambda \leq c_{3,10} t_n^{2HK} n^{-2\beta HK}, \tag{3.23}$$

where the last inequality follows from (2.15).

For  $T_1$ , we use the elementary inequalities  $1 - \cos x \leq x^2$  for every  $x \in \mathbf{R}$  and  $x^\alpha - 1 \leq (x - 1)^\alpha$  for  $x > 1$  and  $0 < \alpha < 1$  to derive that, for all  $s, t \in J_{n,j}$  with  $s < t$ ,

$$\begin{aligned}
T_1 &= \int_{|\lambda| \leq d_{n-1}} |t^{HK} e^{i\lambda \log t} - s^{HK} e^{i\lambda \log s}|^2 f(\lambda) d\lambda \\
&= \int_{|\lambda| \leq d_{n-1}} \left[ (t^{HK} - s^{HK})^2 + 2t^{HK} s^{HK} \left( 1 - \cos \left( \lambda \log \frac{t}{s} \right) \right) \right] f(\lambda) d\lambda \\
&\leq s^{2HK} \left( \left( \frac{t}{s} \right)^{HK} - 1 \right)^2 \int_{\mathbf{R}} f(\lambda) d\lambda + 2t^{2HK} \int_{|\lambda| \leq d_{n-1}} \left( 1 - \cos \left( \lambda \log \frac{t}{s} \right) \right) f(\lambda) d\lambda \\
&\leq s^{2HK} \left( \frac{t}{s} - 1 \right)^{2HK} + 2t^{2HK} \left( \log \frac{t}{s} \right)^2 \int_{|\lambda| \leq d_{n-1}} \lambda^2 f(\lambda) d\lambda \\
&\leq t_n^{2HK} n^{-2\beta HK} + 2t_n^{2HK} n^{-2\beta} c_1 (n-1)^{2\beta(1-HK)} \\
&\leq c_{3,11} t_n^{2HK} n^{-2\beta HK},
\end{aligned} \tag{3.24}$$

where in deriving the last but one inequality, we have used (3.17) and (2.14), respectively.

It follows from (3.22), (3.23), and (3.24) that the  $d$ -diameter of  $J_{n,j}$  satisfies

$$D_j \leq c_{3,12} t_n^{HK} n^{-\beta HK}. \tag{3.25}$$

Hence, similar to (3.21), we use Lemma 2.2 and (3.25) to deduce

$$\mathbf{P}\left\{\max_{s \in J_{n,j}} |X_n^2(s)| > \varepsilon h(t_n)\right\} \leq \exp\left(-c \frac{n^{2\beta HK}}{(\log(n \log n))^{2HK}}\right). \quad (3.26)$$

By (3.16), (3.21), and (3.26), we deduce that for every  $\varepsilon > 0$ ,

$$\begin{aligned} \sum_{n=1}^{\infty} \mathbf{P}\left\{\max_{s \in [0, t_n]} |X_n^2(s)| > \varepsilon h(t_n)\right\} &\leq \sum_{n=1}^{\infty} \sum_{j=0}^{p_n} \mathbf{P}\left\{\max_{s \in J_{n,j}} |X_n^2(s)| > \varepsilon h(t_n)\right\} \\ &\leq c \sum_{n=1}^{\infty} (n^{\beta} \log n) \exp\left(-c \frac{n^{2\beta HK}}{(\log(n \log n))^{2HK}}\right) \\ &< \infty. \end{aligned} \quad (3.27)$$

This proves (3.14) and hence the theorem.  $\square$

By the decomposition of sbfBm and Chung's LIL for the sfBm, we give simple proof of Theorem 3.1.

**Lemma 3.5.** *Let  $S^{H,K}$  be an sbfBm, and assume that  $\{W_t, t \geq 0\}$  is a standard Brownian motion independent of  $S^{H,K}$ . Let  $X^K$  be the process defined by*

$$X_t^K = \int_0^{\infty} (1 - e^{-\theta t}) \theta^{-\frac{1+K}{2}} dW_{\theta}. \quad (3.28)$$

*Then, the processes  $\left\{\sqrt{\frac{K}{\Gamma(1-K)}} X_{t^{2H}}^K + S^{H,K}(t), t \geq 0\right\}$  and  $\{S_t^{HK}, t \geq 0\}$  have the same distribution, where  $\{S_t^{HK}, t \geq 0\}$  is an sfBm with Hurst parameter  $HK$ .*

**Proof.** See the proof of Lemma 2.1 in [26]. For the convenience of readers, we give the proof. By (3.28), we know that  $X^K$  is a centered Gaussian process with covariance

$$\begin{aligned} \mathbf{E}(X_t^K X_s^K) &= \int_0^{\infty} (1 - e^{-\theta t})(1 - e^{-\theta s}) \theta^{-1-K} d\theta \\ &= \int_0^{\infty} (1 - e^{-\theta t}) \theta^{-1-K} d\theta - \int_0^{\infty} (1 - e^{-\theta t}) e^{-\theta s} \theta^{-1-K} d\theta \\ &= \int_0^t \left[ \int_0^{\theta} \theta e^{-\theta u} du \right] \theta^{-1-K} d\theta - \int_0^{\infty} \left[ \int_0^{\theta} \theta e^{-\theta u} du \right] e^{-\theta s} \theta^{-1-K} d\theta \\ &= \int_0^t \left( \int_0^{\infty} \theta^{-K} e^{-\theta u} d\theta \right) du - \int_0^t \left( \int_0^{\infty} \theta^{-K} e^{-\theta(u+s)} d\theta \right) du \\ &= \frac{\Gamma(1-K)}{K} [t^K + s^K - (t+s)^K]. \end{aligned} \quad (3.29)$$

Let  $Y_t = \sqrt{\frac{K}{\Gamma(1-K)}} X_{t^{2H}}^K + S^{H,K}(t)$ . Then, from (1.1) and (3.29), we have, for  $s, t \geq 0$ ,

$$\begin{aligned} \mathbf{E}(Y_s Y_t) &= \frac{K}{\Gamma(1-K)} \mathbf{E}(X_{s^{2H}}^K X_{t^{2H}}^K) + \mathbf{E}(S^{H,K}(s) S^{H,K}(t)) \\ &= t^{2HK} + s^{2HK} - (t^{2H} + s^{2H})^K + (t^{2H} + s^{2H})^K - \frac{1}{2}(t+s)^{2HK} - \frac{1}{2}|t-s|^{2HK} \\ &= t^{2HK} + s^{2HK} - \frac{1}{2}(t+s)^{2HK} - \frac{1}{2}|t-s|^{2HK}, \end{aligned}$$

which completes the proof.  $\square$

Lemma 3.5 implies that

$$\{S_t^{H,K}, t \geq 0\} \stackrel{d}{=} \left\{ S_t^{HK} - \sqrt{\frac{K}{\Gamma(1-K)}} X_{t^{2H}}^K, t \geq 0 \right\}, \quad (3.30)$$

where  $\stackrel{d}{=}$  means equality of all finite dimensional distributions.

**Simple proof of Theorem 3.1.** [22] established Chung's LIL for sfBm  $S^H$ . Namely, there exists a positive and finite constant  $c_{3,13}$  such that

$$\liminf_{r \rightarrow 0} \frac{\max_{t \in [0,r]} |S_t^H|}{r^H / (\log \log(1/r))^H} = c_{3,13}, \quad \text{a.s.}$$

The decomposition (3.30) allows us deduce Chung's LIL for the sbfBm, from the same result for the sfBm with Hurst parameter  $HK$ , with the same constant.

## 4 Applications

In this section, we give some applications of the results in this article. For estimating the self-similar index  $HK$  of a sbfBm. We introduce an estimator for the index  $HK$  of  $S^{H,K}$  given by

$$\hat{HK}_n(p) = \frac{1}{p \log n} \log \frac{S_n(p)}{S_{n^2}(p)},$$

where

$$S_{v_n}(p) = \frac{1}{v_n} \sum_{j=1}^{\lfloor v_n \rfloor} \left| S^{H,K} \left( \frac{j}{v_n} \right) - S^{H,K} \left( \frac{j-1}{v_n} \right) \right|^p.$$

**Theorem 4.1.** For any  $p \geq 1$ , we have  $\hat{HK}_n(p) \rightarrow HK$  almost surely as  $n \rightarrow \infty$ .

**Proof.** By Theorem 2.1, we have, as  $n \rightarrow \infty$ ,

$$\begin{aligned} \hat{HK}_n(p) &= \frac{1}{p \log n} \log \frac{S_n(p)}{S_{n^2}(p)} \\ &= \frac{1}{p \log n} \log \left( \frac{n^{pHK} S_n(p)}{n^{2pHK} S_{n^2}(p)} n^{pHK} \right) \\ &= \frac{1}{p \log n} \log \left( \frac{n^{pHK} S_n(p)}{n^{2pHK} S_{n^2}(p)} \right) + HK \\ &\rightarrow HK, \quad \text{a.s.} \end{aligned}$$

Thus, we finish the proof.  $\square$

**Remark 4.1.** We cannot obtain the estimators of  $H$  and  $K$ , respectively, similar to the estimators of  $H$  and  $K$  for bfBm in [10]. Because the limit in (2.1) has no relation to  $K$ .

In the following, we give some applications of the decomposition of sbfBm.

Recall that a continuous process  $\{X_t, t \in [0, T]\}$  admits  $\alpha$ -variation (resp.  $\alpha$ -strong variation) if the limit in probability

$$\lim_{n \rightarrow \infty} \sum_{i=0}^{n-1} \left| X_{\frac{(i+1)t}{n}} - X_{\frac{it}{n}} \right|^\alpha$$

(resp.

$$\lim_{\varepsilon \rightarrow 0} \frac{1}{\varepsilon} \int_0^t |X_{s+\varepsilon} - X_s|^\alpha ds,$$

exists for every  $t \in [0, T]$ .

Then, we have:

**Theorem 4.2.** *The  $\alpha$ -variation (resp.  $\alpha$ -strong variation) of sbfBm is  $C_{HK}t$ , where  $C_{HK} = \mathbf{E}(|\xi|^{HK})$  and  $\xi$  is a standard normal random variable.*

**Proof.** The results follow easily from (3.29) and the variation in  $X^k$  is 0, since  $X^k$  is absolutely continuous. (Refer also the proofs of Proposition 4 in [27] and Proposition 3.6(a) in [28]).  $\square$

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