

Research Article

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A random von Neumann theorem for uniformly distributed sequences of partitions

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Abstract: In this paper, we prove a theorem that confirms, under a supplementary condition, a conjecture concerning random permutations of sequences of partitions of the unit interval.

Keywords: number theory, discrepancy, uniformly distributed sequences of partitions, probability theory

MSC 2020: 11-xx, 40-xx, 60-xx

1 Introduction

The general study of uniformly distributed sequences of partitions was initiated in [1], inspired by a beautiful construction and result from the study by Kakutani [2]. The subject is closely related to the theory of uniformly distributed sequences, initiated in [3]. There are two classical references for the subject: [4] and [5].

Kakutani took the interval $I = [0, 1]$, a number $\alpha \in]0, 1[$ and divided the interval in proportion $\alpha : 1 - \alpha$. Then, he divided the longest interval of this partition in the same proportion and iterated the procedure dividing always the longest interval of the n th partition, so as to obtain a sequence of partitions of $]0, 1[$. If at a certain step there were two or more intervals of maximal length, they were divided simultaneously.

Kakutani proved that this sequence of partitions (denote it by $\{\alpha^n I\}$) is *uniformly distributed*, which means that if $\alpha^k I = \{0 < t_1^k < t_2^k < \dots < t_{N_k}^k < 1\}$ is the k th partition, then

$$\lim_{k \rightarrow \infty} \frac{1}{N_k} \sum_{i=1}^{N_k} f(t_i^k) = \int_0^1 f(t) dt,$$

for every continuous function f .

In other words, the discrete measure concentrated in the points t_i^k converges weakly to the Lebesgue measure on $[0, 1]$.

The construction has been generalized in [1]. Let ρ be any non trivial finite partition of I .

In the first step, the longest interval(s) of ρ is subdivided positively homothetically to ρ . The partition obtained in this manner is denoted by $\rho^2 I$. In the second step, the same procedure is repeated on the longest interval(s), operating with ρ on $\rho^2 I$. Iteration of this procedure leads to a sequence of partitions $\{\rho^n I\}$.

If $\rho = \{[0, \alpha], [\alpha, 1]\}$, one gets Kakutani's sequence.

The following theorem includes the results of the study by Kakutani ([1], Theorem 2.7)).

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Theorem 1. *The sequence $\{\rho^n I\}$ is uniformly distributed.*

There are interesting connections between the theory of u.d. sequences of partitions and u.d. sequences of points. This connection is far reaching in the construction of a significant subclass of ρ -refinements, the so-called *LS*-sequences. The subject was initiated by the present author in [6], and it is connected with the van der Corput sequences of points.

LS-sequences are constructed starting from the partition ρ_{LS} made of $L + S$ intervals (L and S are positive integers) of length β and β^2 , respectively, where β is the positive solution of the equation $L\beta + S\beta^2 = 1$.

The present paper is concerned with a result in the domain of uniformly distributed sequences of partitions, related to a proposition by von Neumann for uniformly sequences of points [7].

Theorem 2. *If $\{x_n\}$ is a dense sequence of points in $[0, 1]$, then there exists a rearrangement of these points, $\{x_{n_k}\}$, which is uniformly distributed.*

One of the consequences of von Neumann's result is that there are many u.d. sequences of points.

We will now introduce the definitions we need.

Definitions. Given a partition $\pi = \{[t_{i-1}, t_i], 1 \leq i \leq N\}$, we denote by $l_i = t_i - t_{i-1}$ the length of its i th interval. The *diameter* of π , denoted by L , is equal to $\max_{1 \leq i \leq N} l_i$.

Given a sequence of partitions $\{\pi_k\}$, we say that it is *dense* if, denoted by L_k the diameter of π_k , $\lim_{k \rightarrow \infty} L_k = 0$.

If $\pi = \{[t_{i-1}, t_i], 1 \leq i \leq N\}$ is a partition, its *random permutation* is a partition $\pi' = \{[s_{h-1}, s_h], 1 \leq h \leq N\}$ defined by the points $s_h = \sum_{j=0}^h l_{i_j}$, for $0 \leq h \leq N$, where $s_0 = 0$ and the indices $\{i_j\}$, are successively taken at random, with probability $\frac{1}{N}$, from the set $\{i_s : 1 \leq s \leq N\}$.

We will denote by $\pi!$ the set of all the $N!$ permutations of π .

2 Main results

In a previous paper [8], we proved the following result.

Proposition 3. *If $\{\pi_n\}$ is a dense sequence of partitions, then there exists a sequence of partitions $\{\sigma_n\}$, with $\sigma_n \in \pi_n!$, which is uniformly distributed.*

In the same paper, we made the following conjecture.

Conjecture. *If $\{\pi_n\}$ is a dense sequence of partitions and we select at random a partition $\sigma_k \in \pi_n!$, then $\{\sigma_k\}$ is uniformly distributed with probability 1.*

We need some preliminary calculations.

Let $q \in]0, 1[$ and denote by $N_k(q)$ the integer such that

$$\frac{N_k(q)}{N_k} \leq q < \frac{N_k(q) + 1}{N_k}.$$

Consider a sequence $\{q_m\}$ of points, which is dense in $[0, 1]$. For later convenience, we will denote by $N_k(m)$ the integer $N_k(q_m)$.

Select at random from the N_k intervals of π_k , with probability $\frac{1}{N_k}$, $N_k(m)$ intervals. Denote by ξ_i^k the length of the interval selected in the i th draw ($1 \leq i \leq N_k(m)$) and consider the random variable

$$\eta_k^m = \sum_{i=1}^{N_k(m)} \xi_i^k.$$

Obviously,

$$E(\eta_k^m) = \sum_{i=1}^{N_k(m)} E(\xi_i^k) = \sum_{i=1}^{N_k(m)} \frac{1}{N_k} = \frac{N_k(m)}{N_k},$$

hence,

$$|E(\eta_k^m) - q_m| \leq \frac{1}{N_k}.$$

It is easy to see that the second moment of η_k^m is uniformly bounded for any sequence of partitions.

This, together with the independence of the η_k^m 's (for $k \in \mathbb{N}$), would allow us to apply the strong law of large numbers and to conclude that, when k tends to infinity, the sequence η_k^m tends to q_m in the Cesàro mean (and nothing more, at least following this line of thought).

But this is not what we were looking for.

If we want to identify a class for which the conjecture is true, we have to make some assumptions on the sequence $\{\pi_k\}$.

A simple sufficient assumption is expressed as follows:

$$\sum_{i=1}^{\infty} L_k^2 < \infty.$$

Theorem 4. *If the series of squares of diameters of $\{\pi_k\}$ is convergent, then its random permutations σ_k are uniformly distributed with probability 1.*

Proof. We have

$$\begin{aligned} \text{Var}(\eta_k^m) &= E\left(\left(\sum_{i=1}^{N_k(m)} \xi_i^k - \sum_{i=1}^{N_k(m)} \frac{1}{N_k}\right)^2\right) \\ &= E\left(\sum_{i=1}^{N_k(m)} \left(\xi_i^k - \frac{1}{N_k}\right)^2\right) \\ &= E\left(\sum_{i=1}^{N_k(m)} (\xi_i^k)^2 - 2 \sum_{i=1}^{N_k(m)} \xi_i^k \frac{1}{N_k} + \sum_{i=1}^{N_k(m)} \frac{1}{N_k^2}\right) \\ &= E\left(\sum_{i=1}^{N_k(m)} (\xi_i^k)^2\right) - 2 \frac{1}{N_k} E\left(\sum_{i=1}^{N_k(m)} \xi_i^k\right) + E\left(\sum_{i=1}^{N_k(m)} \frac{1}{N_k^2}\right) \\ &\leq E\left(\sum_{i=1}^{N_k} (\xi_i^k)^2\right) - \frac{2}{N_k} E\left(\sum_{i=1}^{N_k} \xi_i^k\right) + \sum_{i=1}^{N_k} \frac{1}{N_k} \\ &= E\left(\sum_{i=1}^{N_k} (\xi_i^k)^2\right) - \sum_{i=1}^{N_k} \frac{1}{N_k} \leq \sum_{i=1}^{N_k} L_k^2 - 1 < \sum_{i=1}^{N_k} L_k^2. \end{aligned}$$

Apply now the Čebišëv inequality. By our assumption, we have, for every $\varepsilon > 0$ (and every m),

$$\sum_{k=1}^{\infty} P(|\eta_k^m - E(\eta_k^m)| > \varepsilon) \leq \sum_{n=1}^{\infty} \frac{\text{Var}(\eta_k^m)}{\varepsilon^2} < \infty.$$

Recalling that $E(\eta_k^m)$ tends to q_m and applying the Borel-Cantelli lemma [9, Theorem 4.2.1], we obtain that

$$\lim_{k \rightarrow \infty} \eta_k^m = q_m$$

almost surely for every $m \in \mathbb{N}$.

The set $\{q_m\}$ is countable; therefore, the aforementioned limit holds almost surely for all the values of m simultaneously.

Observe now that $\lim_{k \rightarrow \infty} \frac{N_k(q)}{N_k}$ is an increasing function of q . Therefore, it follows that, almost surely,

$$\lim_{k \rightarrow \infty} \frac{N_k(q)}{N_k} = q$$

for every $q \in [0, 1]$.

In other words, the empirical distribution function F_k of σ_k tends almost surely to the distribution function of the random variable U uniformly distributed on $[0, 1]$.

On the other hand, convergence in distribution is known to be equivalent to weak convergence, so the desired conclusion follows. \square

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